Adaptive Bayesian Sum of Trees Model for Covariate Dependent Spectral Analysis

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Examples of Biomedical Time Series Data

Clinicians and researchers collect a variety of time series data whose oscillatory patterns are of interest.

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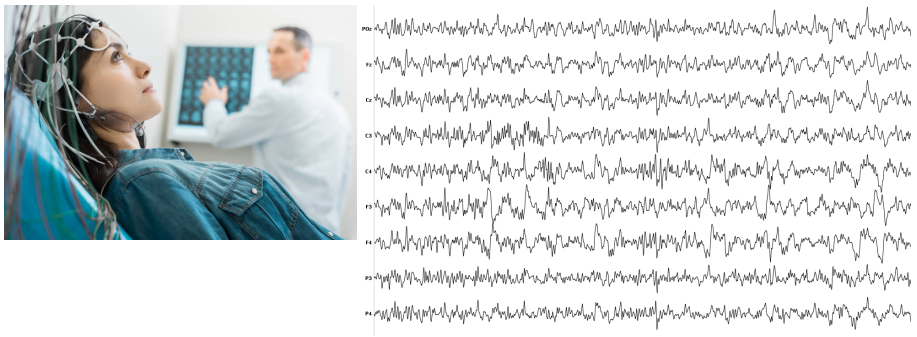
Electroencephalography (EEG)



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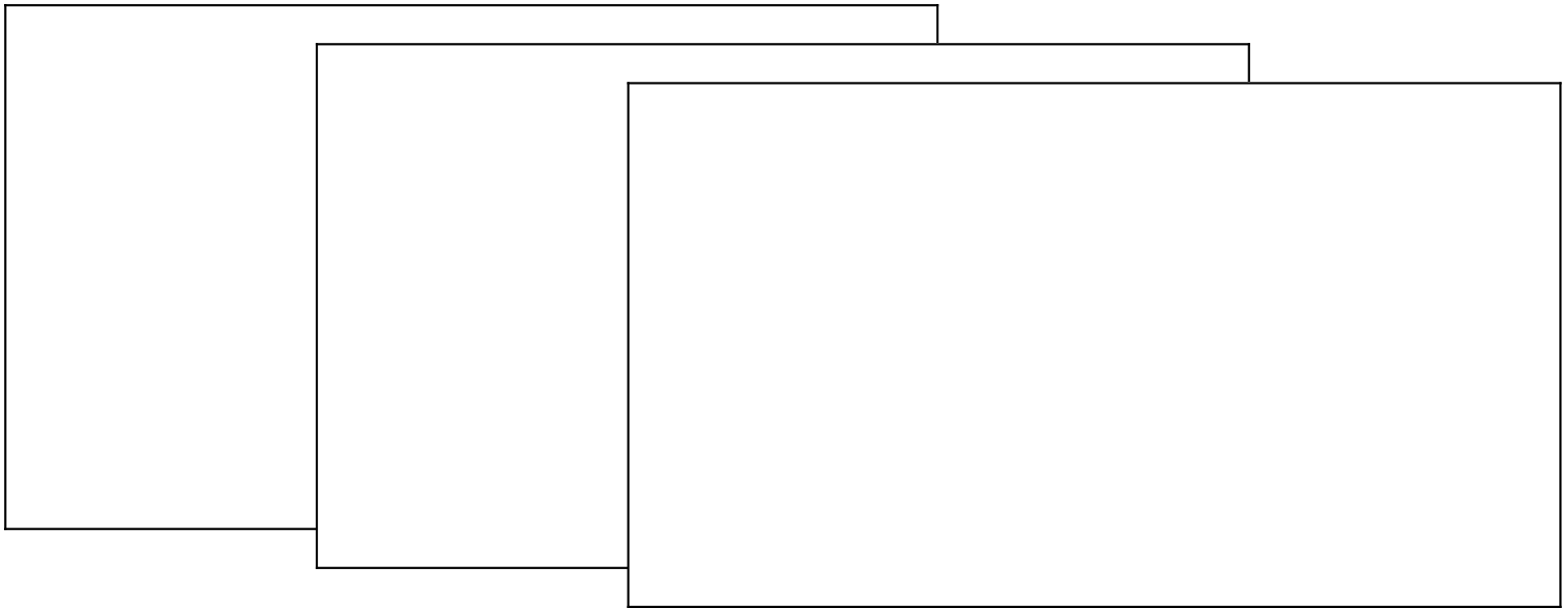

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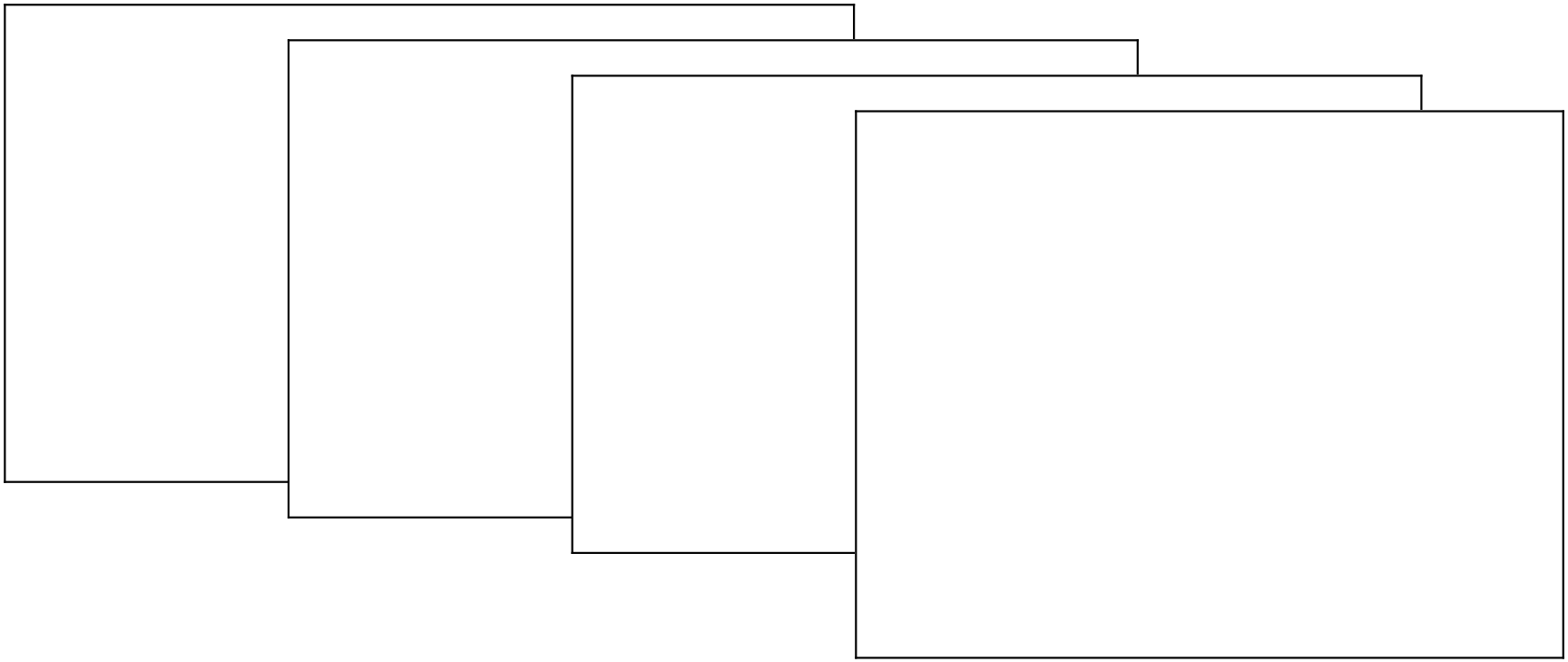

Gait Maturation Study [Hausdor et al. (1999)]

Maturation of gait dynamics

Immature gait in very young children results in unsteady walking patterns and frequent falls.

Gait is relatively mature by age 3. However, neuromuscular control continues to develop well beyond age 3.

Researchers are interested in determining if stride-to-stride dynamics continue to become more steady and regular beyond age 3.

Gait Maturation Study [Hausdor et al. (1999)]

Stride-to-stride time series data

- N = 50 healthy children ages 3-14.
- T = 256 stride times recorded after removing stride times in the rst 60 seconds and last 5 seconds.
- Age, gender, height, weight, leg length, and gait speed are also collected for each child.

Goal: To better understand the maturation of gait dynamics with age in the presence of other related covariates.

Data From Three Subjects

Stationary Time Series

Consider a zero-mean stationary time series X_t.

Cramer Representation [Cramer (1942)]:

$$X_t = \int_{1=2}^{1=2} A(\ ) \exp(2\ i\ t) dZ(\ ):$$

Power spectrum: $f(\ ) = jA(\ )j^2$.

The power spectrum represents a decomposition of variance over frequencies.

$$\text{Var}(X_t) = \int_{1=2}^{1=2} f(\ ) d\ .$$

Two Simulated Examples

Two More Simulated Examples

Periodogram Estimator

Periodogram from $X = (X_1, \ldots, X_T)'$:

$$I(\omega_k) = \frac{1}{T}\left|\sum_{t=1}^{T} X_t \exp(-2\pi i \omega_k t)\right|^2.$$

$\omega_k = k/T$, $k = 1, \ldots, n = \lfloor (T-1)/2 \rfloor$.

Unbiased but noisy estimates of $f(\omega)$.

Approximately distributed as scaled χ^2 to provide the Whittle likelihood:

$$p(x|f) \propto (2\pi)^{-n} \prod_{k=1}^{n} \exp\left\{-\frac{1}{2}[\log f(\omega_k) + I(\omega_k)/f(\omega_k)]\right\}$$

Smoothing

Periodogram can be smoothed to obtain a consistent estimate.

One approach - Bayesian penalized linear spline [Wahba (1990)]:

$$\log f(\ ) \quad + \sum_{s=1}^{S} \ _s \cos(2\ s\ )$$

Priors [Rosen, Wood, and Sto er (2012)]

$$N(0;\ ^2)$$
$$N(0;\ ^2 D_S); \text{where } D_S = \text{diag}(f^{p}\ \overline{2}\ sg\ ^2)$$
$$\text{half-t}$$

Sampling via Metropolis-Hastings and Gibbs steps.

Covariate-dependent Power Spectrum

Covariate-dependent Cramer Representation

$$X_{\ell t} = \int_{1/2}^{1/2} A(\omega_{\ell}; \boldsymbol{\omega}) \exp(2\pi i t \omega) dZ_{\ell}(\omega);$$

where $\boldsymbol{\omega} = (\omega_1, \ldots, \omega_P)^0$ is a P-dimensional covariate vector, and $\ell = 1, \ldots, L$ independent subjects.

Covariate-dependent power spectrum:
$f(\omega; \boldsymbol{\omega}) = |A(\omega; \boldsymbol{\omega})|^2$.

Goal: Develop an adaptive method that can capture both smooth and abrupt changes in power spectra across multiple covariates and provide a tool for variable selection.

One Option: Tree-based Approach

Regression tree illustration

Tree-based models provide a exible and parsimonious approach for partitioning multiple covariates.

For scalar responses, Bayesian Additive Regression Tree (BART) model [Chipman et al. (2010)]

Adaptive Bayesian Sum of Trees Model

Idea: Develop a Bayesian sum-of-trees model for $f(g;)$

$$\log f(!\ ;\ ) \quad \sum_{j=1}^{M} \sum_{b=1}^{B_j} (!\ ; U_j; b) \log f_{bj}( );$$

M is the number of trees

U_j represents the jth tree that has B_j terminal nodes

is a function that identifies terminal node membership such that $(!\ `; U; b) = 1$ if the $`$th observation falls into the bth terminal node and $(!\ `; U; b) = 0$ otherwise.

Model specification for $\log f_{bj}( )$ then follows directly from the Bayesian penalized linear spline introduced previously.

Tree Structure Priors

A regularization prior is applied to encourage each tree to be a weak learner:

$$\Pr(\text{SPLIT}) = \quad (1 + d) \quad ; \qquad 2\ (0; 1); \quad 2\ [0; 1\ );$$

d is the depth of a tree, $= 0:95$ and $= 2$ as default. [Chipman et al. (2010)]

Terminal node parameters and trees are assumed to be independent a priori.

Uniform priors on split variables and cut points.

Sparsity-inducing Dirichlet prior on split variables can also be used for improved variable selection. [Linero, 2018]

Sampling Scheme

Back tting Markov chain Monte Carlo (MCMC) on `residual' of periodogram

$$R_{`j}(\ _k) = \log I_{`}(\ _k) \sum_{i \neq j} \sum_{b=1}^{B_j} (!\ ; U_i ; b) \log f_{bi}(\ )$$

allows for updating each individual tree structure in turn.

Reversible jump MCMC

- Birth: splitting a terminal node into two child nodes
- Death: dropping two terminal child nodes belonging to the same internal node
- Change: modifying the variable and cut point associated with an internal node with two terminal child nodes

Overview of Proposed Approach

Adaptively partition covariate space using tree structures.

Bayesian penalized spline model for local spectra estimation within each terminal node.

Bayesian Back tting MCMC and Reversible jump MCMC techniques to sample from posterior of the trees

Inference averaged over distribution of trees.

Run times

Mean run times for a single tree update over 100 replicates of the three simulation settings with $M = 5$ trees.

Simulated Abrupt+Smooth Example

$$\text{AR}(1): x_{`t} = \quad `x_{`t \quad 1} + \quad `t$$

$$` =$$

$$\begin{cases} 0{:}7 + 1{:}4!_2 & \text{for } 0 \quad !_1 < 0{:}5 \\ 0{:}9 \quad 1{:}8!_2 & \text{for } 0{:}5 \quad !_1 \quad 1; \end{cases}$$

$$` = 1; \quad ; L = 100 \text{ subjects}$$

$$t = 1; \quad ; T = 250$$

$$!_1; !_2 \overset{i:i:d:}{} U(0; 1)$$

$$`_t \overset{i:i:d:}{} N(0; 1)$$

Estimation Accuracy for Abrupt+Smooth Simulation

Simulated Latent Variable Example

AR(2) :

$$x_{\grave{}t} = \quad z_{\grave{}1}x_{\grave{}t \quad 1} + \quad z_{\grave{}2}x_{\grave{}t \quad 2} + \quad _{\grave{}t}$$

Latent variable mapping

$$(\quad z_{\grave{}1};\quad z_{\grave{}2}) = \begin{cases} (1{:}5;\quad 0{:}75); z_{\grave{}} = 1 \\ (\quad 0{:}8; 0); z_{\grave{}} = 2 \\ (\quad 1{:}5;\quad 0{:}75); z_{\grave{}} = 3 \\ (0{:}2; 0); z_{\grave{}} = 4 \end{cases}$$

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$$! _1; ! _2 \overset{i{:}i{:}d{:}}{} U(0; 1)$$

$$_{\grave{}t} \overset{i{:}i{:}d{:}}{} N(0; 1)$$

Simulated Latent Variable Example

Red line: true log power spectra

Gray points: log periodogram ordinates

Blue line: estimated log power spectra using the proposed Bayesian sum of trees model

Green line: estimated log power spectra using the competing smooth model

Gait Maturation Analysis

Data:

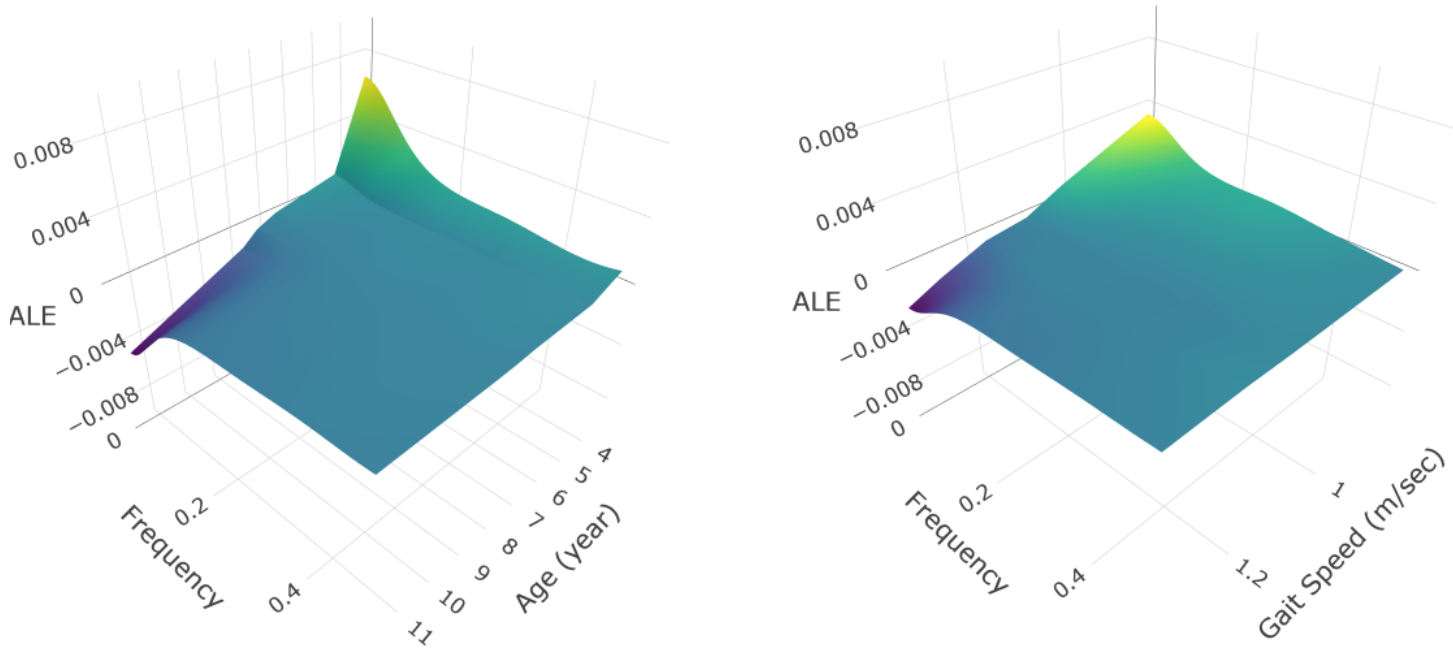
- $N = 50; T = 256$ stride-to-stride time series.
- Ages 3-14 years old.
- Age, gender and gait speed as covariates.

Low frequencies (LF)(0.05-0.25 stride^{-1}) represent uctuations over a longer-term scale (immature gait).

High frequencies (HF) (0.25-0.5 stride^{-1}) represent uctuations over a shorter-term scale (mature gait).

Covariate Effects on Power Spectrum

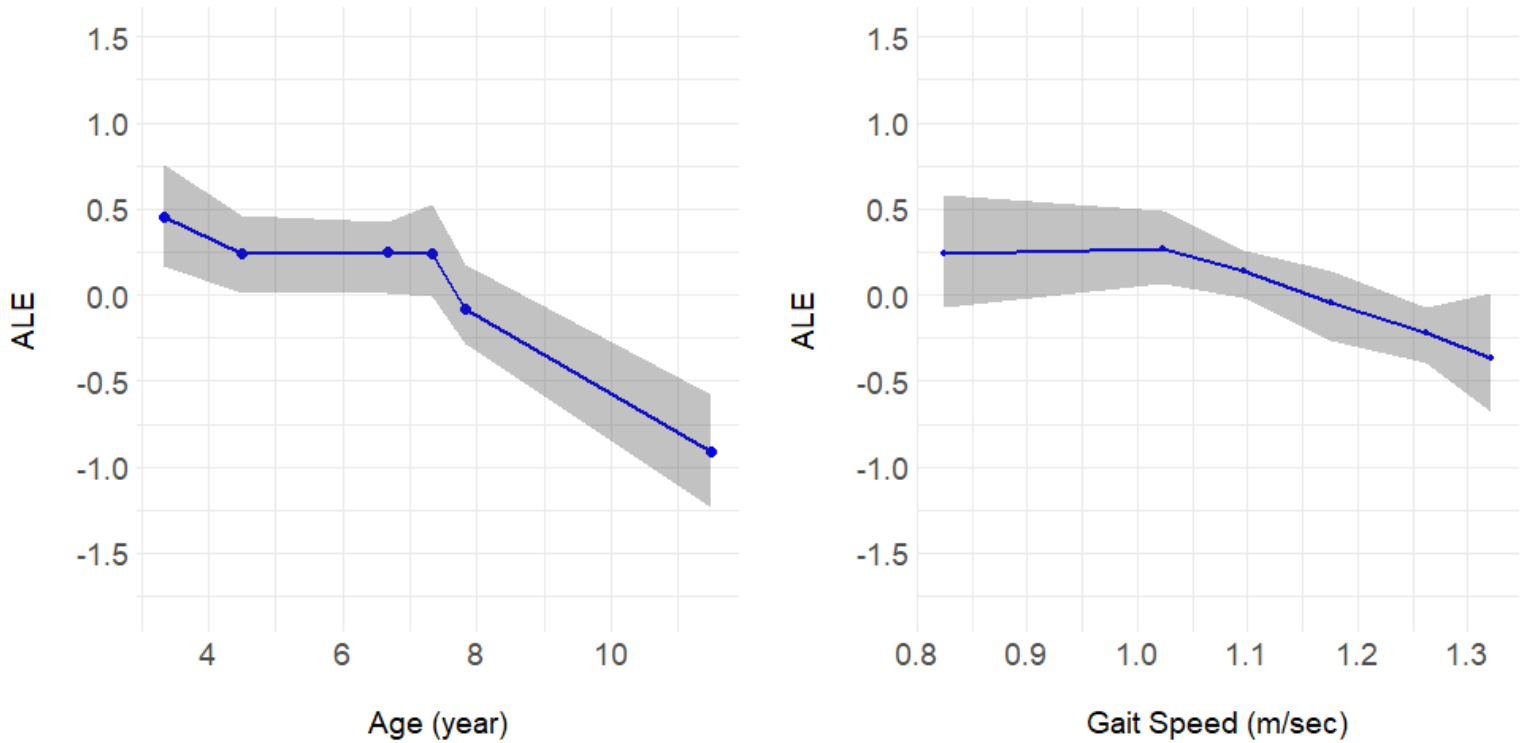
Accumulated local effects (ALE) is a method for evaluating covariate effects.



1. Power over all frequencies decreases as age increase.
2. Power in low frequencies decreases much more with age relative to higher frequencies.

Covariate Effects on Power Spectrum

Low-to-high frequency ratio $\widehat{\frac{\text{LF}}{\text{HF}}}(\boldsymbol{I}) = \frac{\sum_{k \in (0:05;0:25)} \hat{f}(\boldsymbol{I}; k)}{\sum_{k \in [0:25;0:5)} \hat{f}(\boldsymbol{I}; k)}$.



Significant decreases for ages above 7 years and for speeds above 1 m/sec.

Current and Future Work

Current work

- Proposed a nonparametric adaptive Bayesian sum of trees model for covariate-dependent spectral analysis.
- Captures both abrupt and smooth changes.
- Handle complex nonlinear and interaction effects.

Future work

- Extend to time- and covariate-dependent time series.
- Apply alternative partitioning frameworks such as Voronoi tessellations. [Payne et al., 2020]

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THANK YOU!

Gait Maturation Analysis

Accumulated local effects (ALE) is a method for evaluating covariate effects

ALE for $!_j = x$ on the power spectrum is

$$f_{j;\text{ALE}}(x; \ ) = \int_{z_{0;j}}^{x} E_{\boldsymbol{!}_{nj}|!_j} \left[\frac{f(!; \ )}{!_j} \right] !_j = z_j \ dz_j \quad \text{constant}$$

$! = (!_j; !_{nj})$ where $!_j$ denotes the jth covariate and $!_{nj}$ denotes all covariates other than the jth covariate

$\mathbb{Z}_j = \{z_{0;j}; \ldots; z_{H;j}\}$ is a collection of $H + 1$ partition points over the effective support of $!_j$

The constant is a value to vertically center the plot